

# Extending conditional-autoregressive models for space-time disease mapping

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# Acknowledgements

Work in this talk part of a larger collaboration:

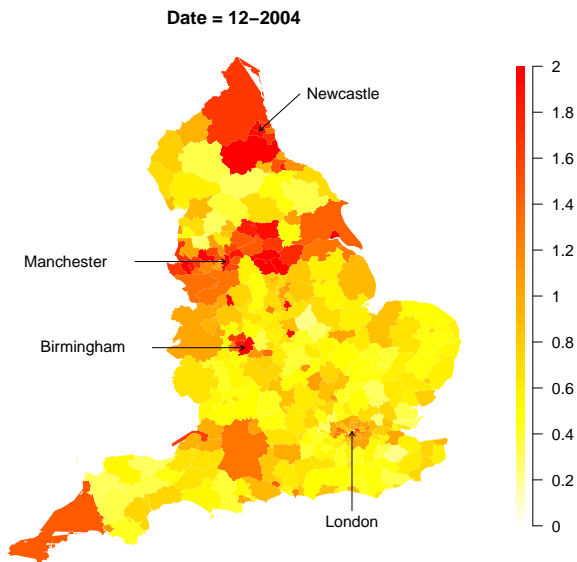
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## Context: England LHAs - Respiratory admission SIR



- ▶ 123 months
- ▶ 323 Local Health Authorities (LHAs)
- ▶ January 2001 to December 2011
- ▶ 42636 data points in total

## Space-time disease mapping: why do it?

- ▶ Understanding if public health policies have an effect over time
- ▶ Understanding health inequality (spatial disease risk)
- ▶ Boundary (spatial discontinuity) detection
- ▶ Clustering of regions with similar (eg. temporal) risk attributes

## Modelling disease counts in space

Typical spatial model might look like

$$Y_i | E_i, R_i \sim \text{Poisson}(E_i R_i)$$

$$\ln(R_i) = \beta_0 + \sum_{j=1}^p x_{ij} \beta_j + \phi_i, \quad i = 1, \dots, N$$

Where

$Y_i$  disease counts

$R_i$

disease risk

$E_i$  expected cases

$\sum_{j=1}^p x_{ij} \beta_j$

covariates

$$\pi(\phi) \propto \exp \left[ -\frac{1}{2\sigma^2} \left( \sum_{i \sim j} w_{ij} (\phi_i - \phi_j)^2 \right) \right]$$

## Localised smoothing

Different strategies achieving this:

Treat  $\phi$ 's prior variance,  $\sigma$ , as spatially varying

- ▶ Brewer and Nolan (2007); Reich and Hodges (2008)

Treat the weights  $w_{ij}$  as random variables

- ▶ Ma, Carlin and Banerjee (2010); Lee and Mitchell (2013)

Use a clustering or grouping prior for the random effects

- ▶ Richardson and Green (2002); Knorr-Held and Raßer (2000)

## Ma, Carlin and Banerjee (2010)

Starting with Poisson model and Intrinsic CAR for random effects:

$$Y_i | E_i, R_i \sim \text{Poisson}(E_i R_i)$$

$$\log(R_i) = \beta_0 + \sum_{j=1}^p x_{ij} \beta_j + \phi_i, \quad i = 1, \dots, N$$

$$p(\phi | \sigma, \{w_{ij}\}) = C(\sigma, \{w_{ij}\}) \exp \left( -\frac{1}{2\sigma^2} \sum_{i \sim j} w_{ij} (\phi_i - \phi_j)^2 \right)$$

Normalising term  $C(\sigma, \{w_{ij}\})$  included to emphasise that  $w_{ij}$  is now being treated as unknown.

## Ma, Carlin and Banerjee (2010); CAR2 model

**Idea:** Call the non-zero  $w_{ij}$ ,  $w_{ij}^+$ , and treat as Bernoulli random variables with unknown probabilities,  $p_{ij}$ .

Transform and smooth  $p_{ij}$  using a further CAR prior.

$$w_{ij}^+ | p_{ij} \sim \text{Bernoulli}(p_{ij}) \text{ and } \text{logit}(p_{ij}) = \mathbf{z}_{ij}'\boldsymbol{\gamma} + \theta_{ij}$$
$$p(\boldsymbol{\theta} | \zeta) \propto \exp\left(-\frac{1}{2\zeta^2} \sum_{ij \sim kl} (\theta_{ij} - \theta_{kl})^2\right)$$

**Pros:** Maintains binary nature of  $w_{ij}$

**Cons:** Assumes relatively strong smoothness over the  $w_{ij}$ ; a lot of parameters to estimate  $\boldsymbol{\phi}$ ,  $w_{ij}$ ,  $\boldsymbol{\theta}$  that can be hard to identify.



## A localised model in space and time

Poisson model for counts in space and time (indexed by  $i$  and  $j$ , respectively)

$$Y_{ij} | E_{ij}, R_{ij} \sim \text{Poisson}(E_{ij} R_{ij})$$
$$\log(R_{ij}) = \beta_0 + \phi_{ij}, \quad i = 1, \dots, N \text{ and } j = 1, \dots, T$$

Letting  $\phi_t = (\phi_{1t}, \dots, \phi_{Nt})$ , prior for spatial random effects is

$$p(\phi_t | \sigma, \{w_{ij}\}) = C(\sigma, \{w_{ij}\}) \exp \left( -\frac{1}{2\sigma^2} \sum_{i \sim j} w_{ij} (\phi_{it} - \phi_{jt})^2 \right)$$

## A localised model in space and time

**Idea:** Treat the non-zero elements of  $w_{ij}$ ,  $w_{ij}^+$ , as random variables on  $[0, 1]$ .

Then transform using  $\text{logit}(w_{ij}^+) = z_{ij}$  and smooth these using a Leroux prior.

$$p(z_{ij}) \propto \exp \left[ -\frac{1}{2\tau^2} \left( \rho \sum_{ij \sim kl} (z_{ij} - z_{kl})^2 + (1 - \rho) \sum_{ij} (z_{ij})^2 \right) \right]$$

**Pros:** Treating  $w_{ij}$  as non-binary makes things slightly easier; Leroux prior is a bit more flexible than the Intrinsic CAR model.

**Cons:** Still a lot of parameters to estimate

# A localised model in space and time - computation

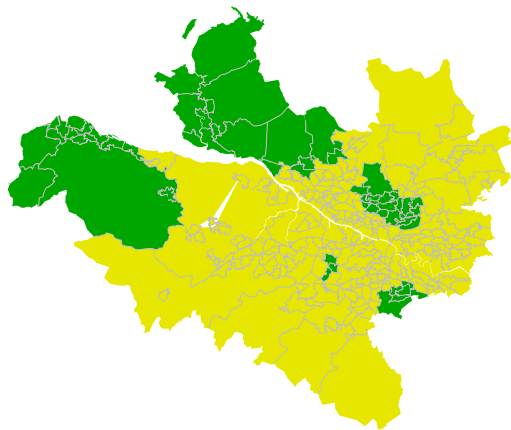
## Need to update $\phi_{N \times T}$

- ▶ Can be done very quickly using one-at-a-time updating in C++ and exploiting matrix sparsity

## Need to update $z_{ij}$

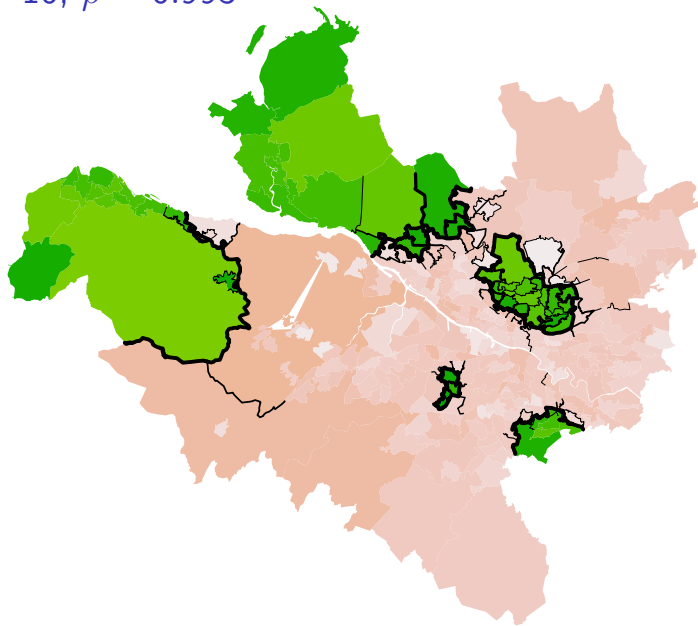
- ▶ Block updating using each blocks conditional prior variance
- ▶ Main bottleneck: recalculating determinant of  $\phi$  prior precision for each block proposal
- ▶ 85% of MCMC time even with sparse Cholesky routine

## Some illustrative simulations

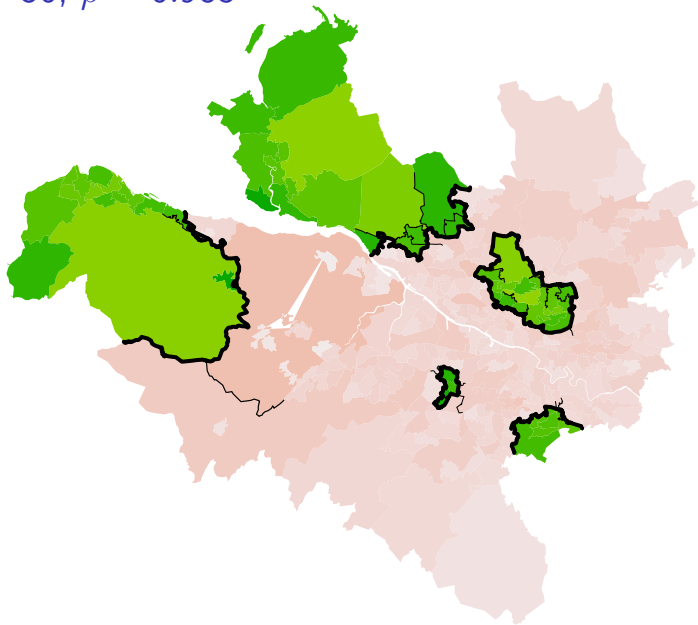


- ▶ Using Glasgow intermediate geographies as a spatial template
- ▶ To investigate effect of increasing time points, generate 10 or 50 repeated realisations of this surface
- ▶ Simulate spatial counts step-function using Poisson with mean of either 100 or 200.
- ▶ Interested in effect of  $\rho$ : fix at 0, or treat as unknown.

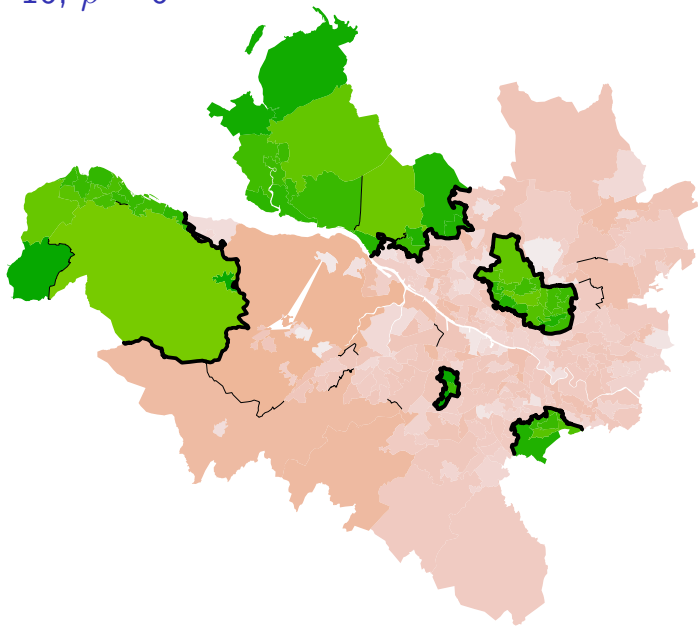
$T = 10; \rho = 0.993$



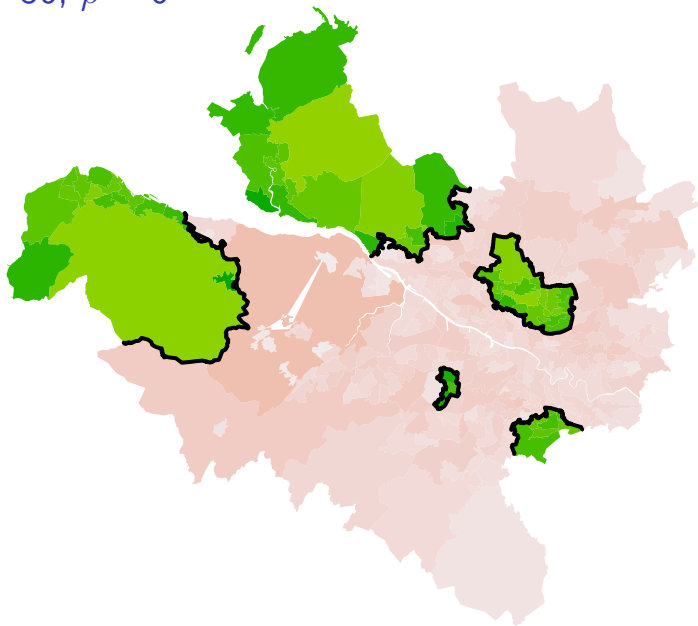
$T = 50; \rho = 0.983$



$T = 10; \rho = 0$



$T = 50; \rho = 0$





## DIC and parameter estimates for each model

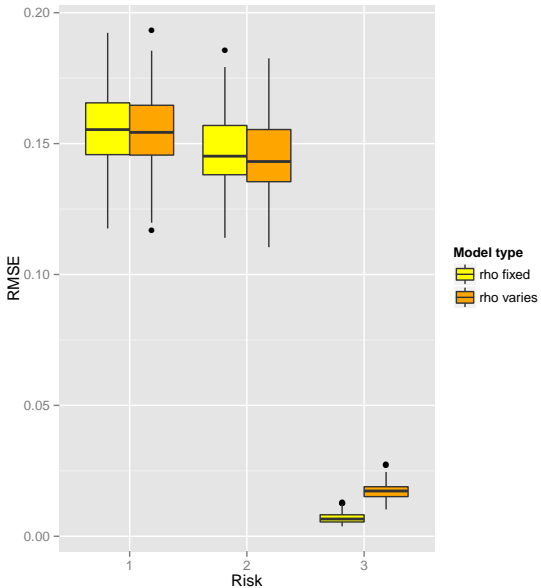
**DIC table for  $T = 10$**

$\rho$	DIC
0.993	25093.07
0	25331.94

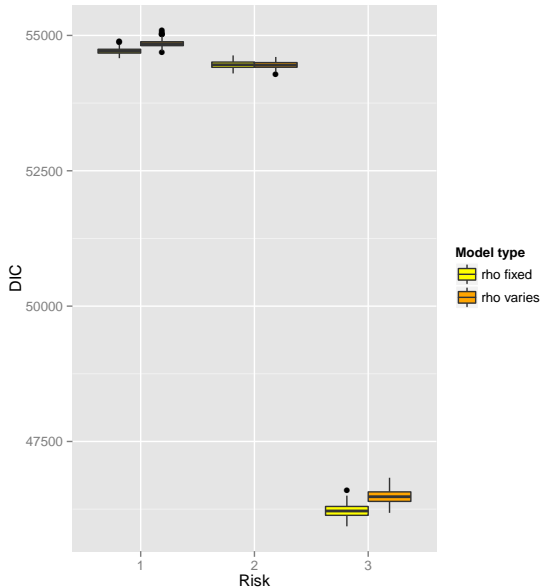
**DIC table for  $T = 50$**

$\rho$	DIC
0.983	127717.5
0	128425.7

## Some more (simulation) results

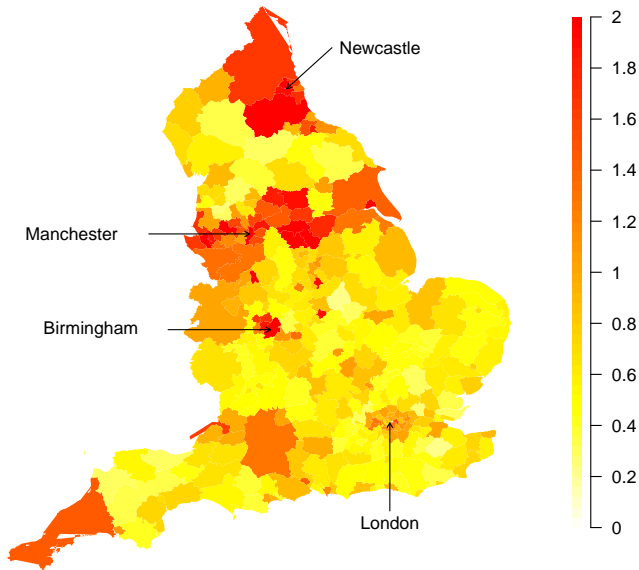


# Some more (simulation) results



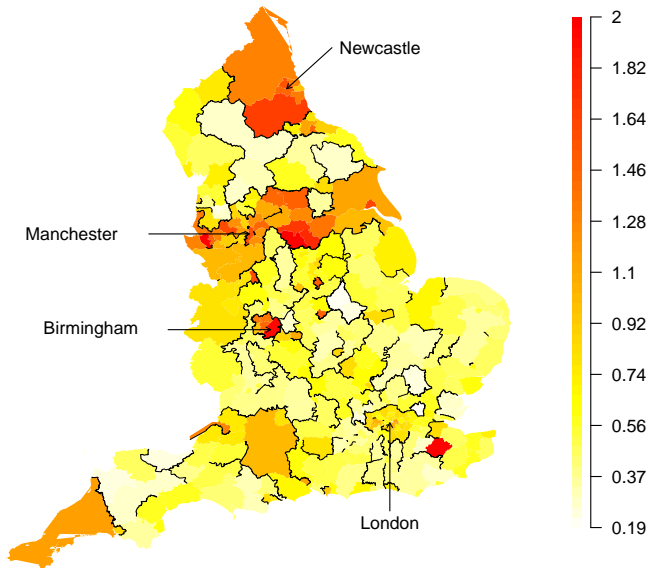
# Application to England LHAs - Data

Date = 12-2004



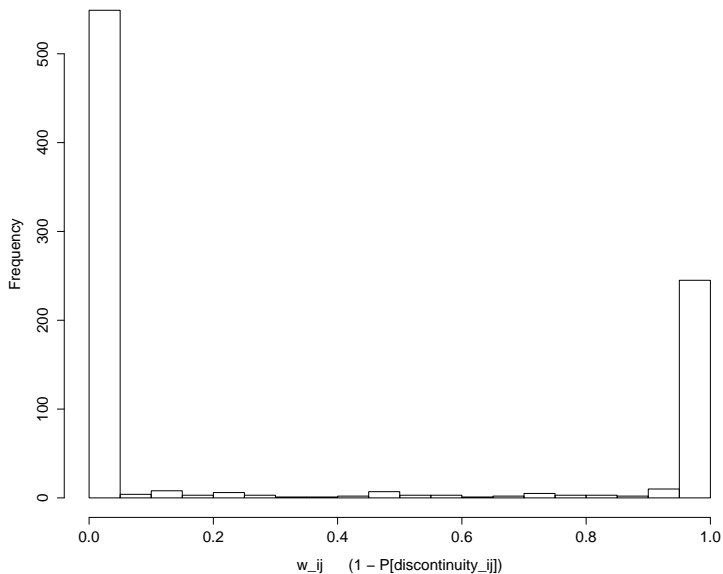
# Application to England LHAs - Results

Date = 12-2004



# Application to England LHAs - Results

Distribution of posterior mean  $w_{ij}$



## Current work

Model is obviously overparameterised

- ▶ Set of spatial random effects for each time and region
- ▶ Doesn't take account of temporal structure

Letting  $\tilde{\phi}_t = (\phi_{t1}, \dots, \phi_{tN})$ , where  $t = 1, \dots, T$ . Assume a model of the form

$$f(\tilde{\phi}_1, \dots, \tilde{\phi}_T) = f(\tilde{\phi}_1) \prod_{t=2}^T f(\tilde{\phi}_t | \tilde{\phi}_{t-1}),$$

where

$$\begin{aligned}\tilde{\phi}_1 &\sim ICAR(\mathbf{0}, \sigma^2 Q^{-1}) \\ \tilde{\phi}_t | \tilde{\phi}_{t-1} &\sim ICAR(\alpha \tilde{\phi}_{t-1}, \sigma^2 Q^{-1}).\end{aligned}$$

## Comments

More work required to understand properties of these models

- ▶ Discontinuity detection rates under realistic surfaces
- ▶ How many replications make this analysis worthwhile?
- ▶ Simulations in progress...

Difficult to incorporate information about likely discontinuity structure into prior - compromise sometimes is required due to computational constraints:

- ▶ This is the reason CAR models are so nice: (matrix) sparsity; well established tricks for model-fitting; easy to write very efficient code.
- ▶ Knorr-Held and Raßer (2000) use a nice clustering prior that can induce smoothness; but requires RJMCMC.



## References

Besag, J., J. York, and A. Mollie (1991). Bayesian image restoration with two applications in spatial statistics. *Annals of the Institute of Statistics and Mathematics*

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Thanks for listening!

