

Gaussian Markov Random Field models for capturing localized residual spatio-temporal confounding in air pollution and health studies

Alastair Rushworth



**International Conference of the Royal Statistical Society,
Sheffield**

2nd September 2014

Acknowledgements

Work in this talk part of a larger collaboration:

University of Glasgow { Duncan Lee
Richard Mitchell }

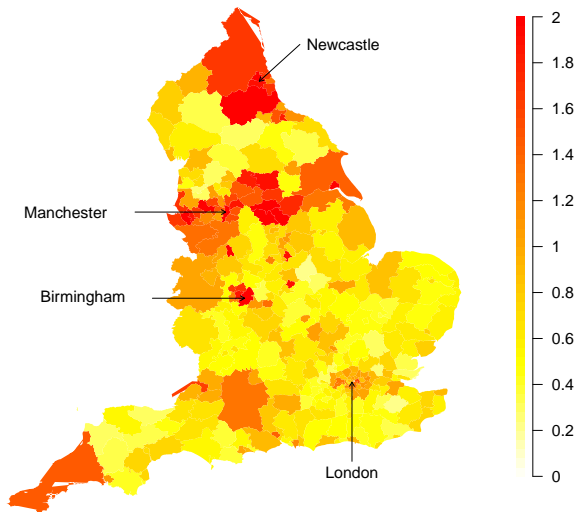
University of Southampton { Sujit Sahu
Sabyasachi Mukhopadhyay }

UK Met Office { Paul Agnew
Christophe Sarran }

Funded by the EPSRC { 
Engineering and Physical Sciences
Research Council }

Context: England LHAs - Respiratory admission SIR

Date = 12-2004



- ▶ 123 months
- ▶ 323 Local Health Authorities (LHAs)
- ▶ January 2001 to December 2011
- ▶ 42636 data points in total

Modelling disease counts in space

Typical spatial model might look like

$$Y_i | E_i, R_i \sim \text{Poisson}(E_i R_i)$$

$$\ln(R_i) = \beta_0 + \sum_{j=1}^p x_{ij} \beta_j + \phi_i, \quad i = 1, \dots, N$$

Where

Y_i disease counts

R_i

disease risk

E_i expected cases

$\sum_{j=1}^p x_{ij} \beta_j$

covariates

$$\pi(\phi) \propto \exp \left[-\frac{1}{2\sigma^2} \left(\sum_{i \sim j} w_{ij} (\phi_i - \phi_j)^2 \right) \right]$$

Localised smoothing

Different strategies achieving this:

Treat ϕ 's prior variance, σ , as spatially varying

- ▶ Brewer and Nolan (2007); Reich and Hodges (2008)

Treat the weights w_{ij} as random variables

- ▶ Ma, Carlin and Banerjee (2010); Lee and Mitchell (2013)

Use a clustering or grouping prior for the random effects

- ▶ Richardson and Green (2002); Knorr-Held and Raßer (2000)

CAR2 model

Ma, Carlin and Banerjee (2010)

Starting with Poisson model and Intrinsic CAR for random effects:

$$Y_i | E_i, R_i \sim \text{Poisson}(E_i R_i)$$

$$\log(R_i) = \beta_0 + \sum_{j=1}^p x_{ij} \beta_j + \phi_i, \quad i = 1, \dots, N$$

$$p(\phi | \sigma, \{w_{ij}\}) = C(\sigma, \{w_{ij}\}) \exp \left(-\frac{1}{2\sigma^2} \sum_{i \sim j} w_{ij} (\phi_i - \phi_j)^2 \right)$$

Normalising term $C(\sigma, \{w_{ij}\})$ included to emphasise that w_{ij} is now being treated as unknown.

CAR2 model

Ma, Carlin and Banerjee (2010)

Idea: Call the non-zero w_{ij} , w_{ij}^+ , and treat as Bernoulli random variables with unknown probabilities, p_{ij} .

Transform and smooth p_{ij} using a further CAR prior.

$$w_{ij}^+ | p_{ij} \sim \text{Bernoulli}(p_{ij}) \text{ and } \text{logit}(p_{ij}) = \mathbf{z}'_{ij}\boldsymbol{\gamma} + \theta_{ij}$$
$$p(\boldsymbol{\theta} | \zeta) \propto \exp\left(-\frac{1}{2\zeta^2} \sum_{ij \sim kl} (\theta_{ij} - \theta_{kl})^2\right)$$

Pros: Maintains binary nature of w_{ij}

Cons: Assumes relatively strong smoothness over the w_{ij} ; a lot of parameters to estimate $\boldsymbol{\phi}$, w_{ij} , $\boldsymbol{\theta}$ that can be hard to identify.

A localised model in space and time

Poisson model for counts in space and time (indexed by i and j , respectively)

$$Y_{ij} | E_{ij}, R_{ij} \sim \text{Poisson}(E_{ij} R_{ij})$$
$$\log(R_{ij}) = \beta_0 + \phi_{ij}, \quad i = 1, \dots, N \text{ and } j = 1, \dots, T$$

Letting $\tilde{\phi}_t = (\phi_{1t}, \dots, \phi_{Nt})$, where $t = 1, \dots, T$.

$$f(\tilde{\phi}_1, \dots, \tilde{\phi}_T) = f(\tilde{\phi}_1) \prod_{t=2}^T f(\tilde{\phi}_t | \tilde{\phi}_{t-1}),$$

where

$$\tilde{\phi}_1 \sim \text{N}(\mathbf{0}, \sigma^2 Q(W, \epsilon))^{-1}$$
$$\tilde{\phi}_t | \tilde{\phi}_{t-1} \sim \text{N}(\alpha \tilde{\phi}_{t-1}, \sigma^2 Q(W, \epsilon)^{-1}).$$

A localised model in space and time

Idea: Treat the non-zero elements of w_{ij} , w_{ij}^+ , as random variables on $[0, 1]$.

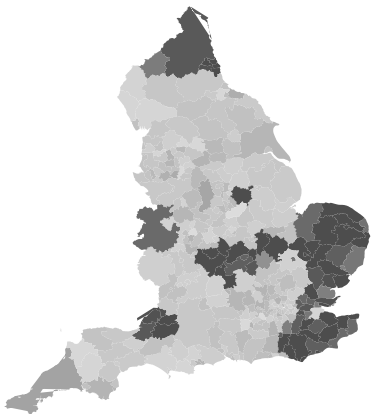
Then transform using $\text{logit}(w_{ij}^+) = z_{ij}$ and smooth these using a Leroux prior.

$$p(z_{ij}) \propto \exp \left[-\frac{1}{2\tau^2} \left(\rho \sum_{ij \sim kl} (z_{ij} - z_{kl})^2 + (1 - \rho) \sum_{ij} (z_{ij})^2 \right) \right]$$

Pros: Treating w_{ij} as non-binary makes things slightly easier; Leroux prior is a bit more flexible than the Intrinsic CAR model.

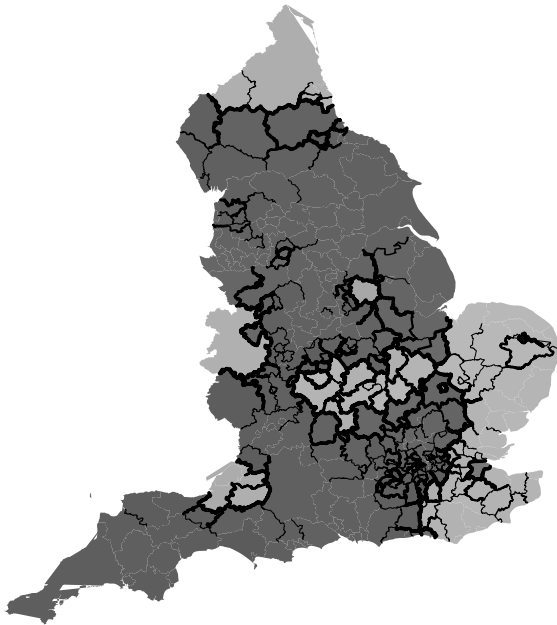
Cons: Still a lot of parameters to estimate

Illustrative simulations

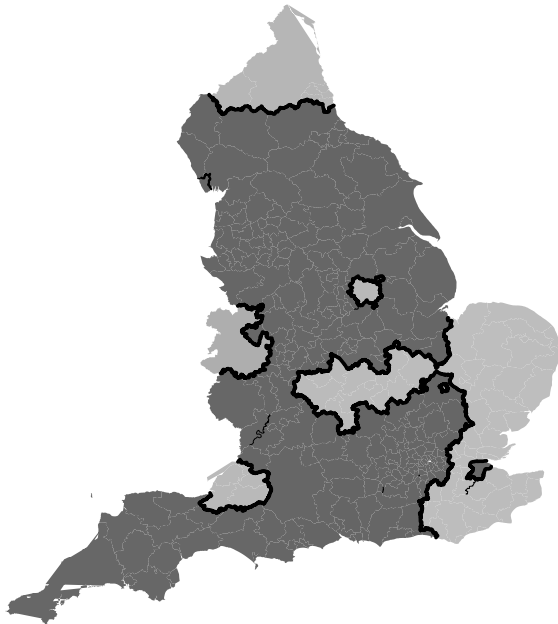


- ▶ English local authorities used as a spatial template
- ▶ What happens to boundary identification with increased temporal replication?
- ▶ What effect does 'smoothing' boundaries have? Can ρ be fixed?

Correlated boundaries: $T = 10$; $\rho = 0.993$



Uncorrelated boundaries: $T = 10$; $\rho = 0$

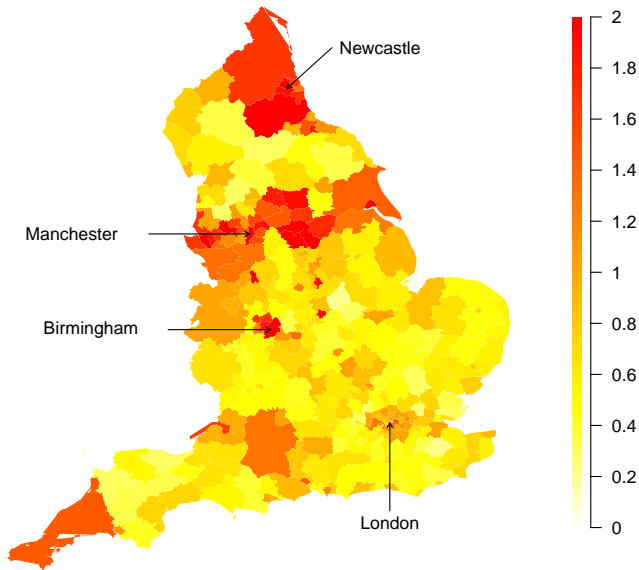


Validation of adaptive model by simulation

| | | RMSE | | |
|-------------------------|-----------|-----------------|------------|------------------|
| | | non adaptive | $\rho = 0$ | ρ varies |
| Temporal replication | Time = 1 | 0.1143 | 0.0971 | 0.0728 |
| | Time = 5 | 0.0865 | 0.0499 | 0.0525 |
| | Time = 20 | 0.0646 | 0.0390 | 0.0405 |
| Relative risk | a = 1 | 0.0363 | 0.0386 | 0.0397 |
| | a = 2 | 0.0868 | 0.0497 | 0.0524 |
| | a = 3 | 0.1085 | 0.0587 | 0.0611 |
| Expected cases | E = 25 | 0.1253 | 0.0686 | 0.0726 |
| | E = 75 | 0.0866 | 0.0495 | 0.0521 |
| | E = 200 | 0.0613 | 0.0375 | 0.0391 |

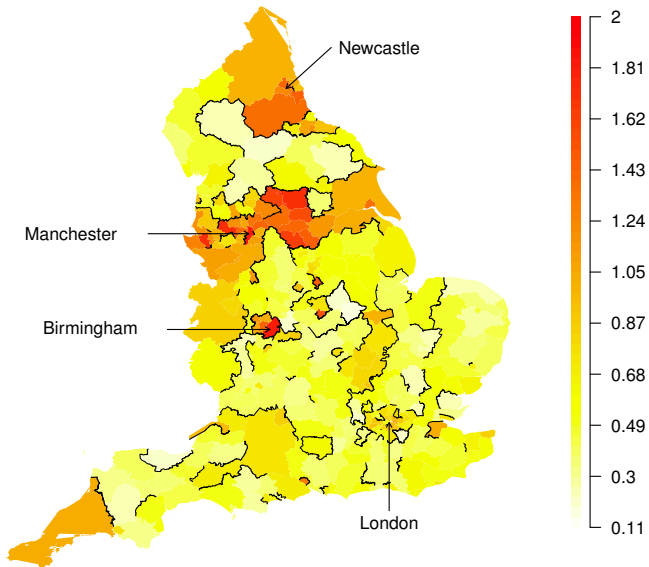
Application to England LHAs - Data

Date = 12-2004



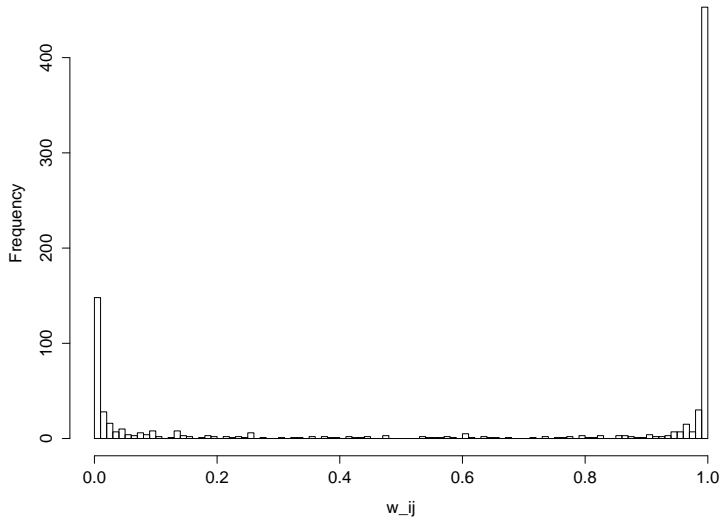
Application to England LHAs - Results

October 2001



Application to England LHAs - Results

Posterior medians of adjacency elements, w_{ij}



Comments

Adaptive model is succesful at identifying boundaries

- ▶ Simulation shows that performance is generally much better than for a standard model
- ▶ The prior distribution for boundaries is attractive as it encourages $w_{ij} = 1$ or $w_{ij} = 0$.
- ▶ Computation is important, and this model exploits matrix sparsity for speed.

In general, difficult to incorporate information about likely discontinuity structure into prior - compromise sometimes is required due to computational constraints:

- ▶ This is why CAR priors are nice: (matrix) sparsity; well established tricks for model-fitting; easy to write very efficient code.
- ▶ Example of alternative: Knorr-Held and Raßer (2000) use clustering prior that can induce smoothness; but requires RJMCMC.

Thanks for listening!

Some references

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