# Spatio-temporal modelling of respiratory health in London

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September 12, 2013

# Acknowledgements

Work is joint with

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- Richard Mitchell University of Glasgow

Funded by the  $\ensuremath{\mathsf{EPSRC}}$ 



Engineering and Physical Sciences Research Council

# Background

- Importance: air pollution is well known to have a negative impact on human health and is still an important public health issue.
- London has a particularly rich history of issues with air pollution and it's subsequent effects. eg. 'Pea-soupers' and the 'Great smog'.
- Estimated 4000 additional deaths due to poor air quality alone in London (Miller, 2010).
- Difficult to unpick the different contributors to ill health without detailed data for risk factors and pollution exposure.

## Goals

- Quantify the impacts of pollutants on respiratory health in Greater London using an ecological design and data at the small area level
- Investigate the effects of different pollutants, and composite indicators.
- Adequately account for unmeasured confounding
- Achieve some level of computational ease important if models are to be widely adopted.

## Talk structure

Introducing the London data

Importance of spatio-temporal modelling

Results

Conclusions and discussion.

## Data

- London has 624 (non-overlapping) electoral wards for which data are available for the period spanning 2002 to 2009 inclusive.
- Health data are available as annualised total of respiratory hospital admissions for each of the areal units.
- Modelled air pollution maps available from DEFRA (www.uk-air.defra.gov.uk)
- Covariate data Average house prices (Price) and proportion of people claiming jobseekers allowance (JSA)

Both  $\mathrm{JSA}$  and  $\mathrm{Price}$  are proxy measures for income and housing deprivation.

## Data summaries

 Spatially averaged respiratory admission and pollutant concentrations (µgm<sup>-3</sup>) between 2002 and 2008.

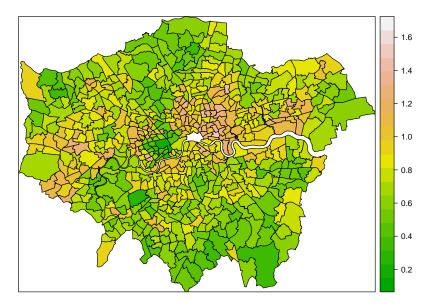
	2002	2003	2004	2005	2006	2007	2008
Resp.	117.00	126.00	132.00	143.00	141.00	144.00	151.00
$PM_{10}$	17.60	20.20	24.60	23.40	22.70	23.80	20.30
$PM_{2.5}$	11.50	17.10	16.70	14.90	15.10	13.50	14.10
CO	372.00	372.00	343.00	338.00	264.00	251.00	229.00
$NO_2$	34.80	36.70	31.90	33.70	32.40	34.40	30.30
$NO_X$	59.50	62.00	53.70	56.30	53.10	57.80	50.10
SO <sub>2</sub>	3.75	6.02	3.10	3.02	3.08	3.16	2.37

As an exploratory measure we plot the Standardised Incidence Ratio (SIR) for each areal unit i, where

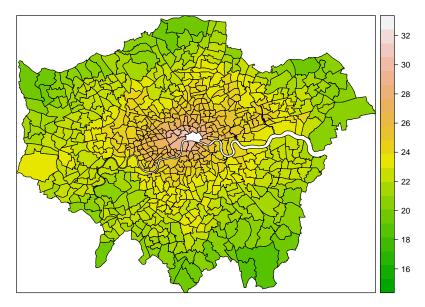
 $SIR_i = \frac{Observed Incidence_i}{Expected Incidence_i}$ 

Where the expected number of cases are calculated using external standardisation based on the population, gender and age distribution of each area.

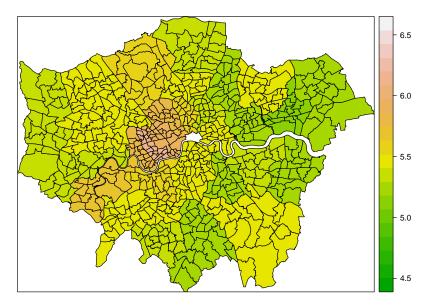
# SIR - London respiratory admissions data (2001)



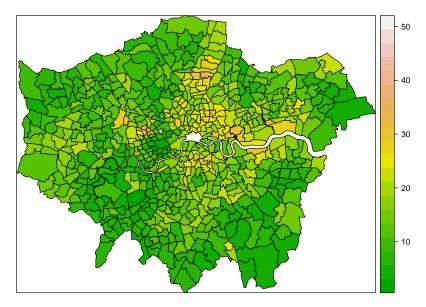
# PM<sub>10</sub> - London air pollution data (2001)



# London house prices (2001)



# Job seekers allowance (2001)



## Residual structure - GLM

Fitting the simple model for counts  $Y_{ij}$  using GLM in R....  $Y_{ij}|E_{ij}, R_{ij} \sim \text{Poisson}(E_{ij}R_{ij})$  $\ln(R_{ij}) = \alpha + \text{JSM}_{ij}\beta_1 + \text{Price}\beta_2 + \text{Poll}_{ij}\beta_3$ 

...but residuals are spatially correlated...



	Morans I	p-value
GLM	0.44	0.00

# Models for spatial confounding

Introduce a set of spatially smooth random effects  $\phi$  following a Gaussian Markov Random Field (GMRF) prior, options include:

- Intrinsic Autoregressive prior (Besag et al. (1991)); assumes strong smoothness of random effects.
- Besag-York-Mollie (Besag et al. (1991)); 2 sets of random effects (indep + spatial).
- Leroux prior (Leroux et al. (1999)); both indep and intrinsic models are special cases.

...and others. Also possible to use a geostatistical model, splines or other smoothers.

## Models for spatial confounding

Leroux prior in space

#### Overall model is now

$$\begin{array}{lll} Y_{ij}|E_{ij},R_{ij} & \sim & \mathrm{Poisson}(E_{ij}R_{ij}) \\ & & \ln(R_{ij}) & = & \alpha + \mathrm{JSM}_{ij}\beta_1 + \mathrm{Price}_{ij}\beta_2 + \mathrm{PM}_{10}\beta_3 + \phi_i \end{array}$$

Where  $\boldsymbol{\phi} = (\phi_1, \dots, \phi_p)$  and

$$\boldsymbol{\phi} \sim N\left(\mathbf{0}, \sigma^2 \left[\rho W^* + (1-\rho)I\right]^{-1}\right)$$

Where  $W^*$  is  $diag(W\mathbf{1}^T) - W$  and W is the adjacency matrix.

Need some additional machinery to allow the random effects dependence in time as well as space.

## Other models for residual spatio-temporal structure

Many exist, although mainly used in settings where no covariates available.

- Bernardinelli et al. (1995) spatially varying slopes and intercepts describing linear temporal changes
- MacNab and Dean (2001) spatially varying spline components to describe non-linear temporal changes.
- Knorr-Held (2000) studies different forms of space time interaction
- ▶ Ugarte. et al (2012) P-spline ANOVA approach
- Lawson. et al (2012) Bayesian mixture, some emphasis on clustering of temporal patterns

### Models for spatio-temporal confounding

Leroux prior in space and time

Let 
$$\tilde{\phi}_t = (\phi_{t1}, \dots, \phi_{tN})$$
. Assume a model of the form  $au$ 

$$f( ilde{\phi}_1,\ldots, ilde{\phi}_t) = f( ilde{\phi}_1)\prod_{t=2}^{\prime}f( ilde{\phi}_t| ilde{\phi}_{t-1})$$

Desirable to allowing the prior to allow temporal independence and strong dependance as special cases. One approach would be

$$f(\tilde{\phi}_1) \sim N\left(\mathbf{0}, \sigma^2 \left[\rho W^* + (1-\rho)I\right]^{-1}\right)$$
  
$$f(\tilde{\phi}_t | \tilde{\phi}_{t-1}) \sim N\left(\alpha \tilde{\phi}_{t-1}, \sigma^2 \left[\rho W^* + (1-\rho)I\right]^{-1}\right)$$

where  $\alpha \in [0,1]$  captures the strength of temporal dependence.

# Prior distributions and inference

We assume

$$\begin{array}{rcl} \rho, \alpha & \sim & U[0,1] \\ \sigma^2 & \sim & U[0,1000] \\ \alpha, \beta_1, \dots, \beta_p & \sim & \mathsf{N}(0,100) \end{array}$$

- Samples from the marginal posterior of α can be drawn using Gibbs sampling
- Metropolis-Hastings is used for  $\beta$ ,  $\rho$ ,  $\sigma^2$ ,  $\phi$ .
- For computational speed, the 624 × 7 vector φ is updated using C++

# Results - is a temporal component needed for London data?

To do this compare 3 scenarios under the proposed spatio-temporal model:

$$\alpha = \begin{cases} 0 & \text{no temporal dependence} \\ 1 & \text{strong temporal dependence} \\ \in [0, 1] & \text{something in between} \end{cases}$$

	DIC	Morans I	p-value
$\alpha = 1$	36125.6	0.010	0.1490
$\alpha = 0$	36986.9	-0.043	0.0000
lpha= 0.85	36074.1	0.019	0.0056

 $\label{eq:Table:DIC} \ensuremath{\mathsf{Table}}\xspace: \mathsf{DIC} \ensuremath{\mathsf{and}}\xspace \ensuremath{\mathsf{residual}}\xspace \ensuremath{\mathsf{correlation}}\xspace \ensuremath{\mathsf{and}}\xspace \ensuremath{\mathsf{residual}}\xspace \ensurema$ 

It might be expected that the overall composition of the air is what increases ill health of a population.

As a simple measure take the mean of each pollutant variable as a composite measure at each site and time. [Each pollutant standardised first so that they have mean 0 and variance 1]

$$\text{Composite}_{ij} = \frac{\sum_{k=1}^{m} \text{Pollutant}_{ijk}}{k}$$

## Effects of individual pollutants and composite measure

	RR	95% CI
PM <sub>10</sub>	1.022	(1.002,1.04)
PM <sub>2.5</sub>	1.032	(1.015, 1.051)
CO	1.023	(1, 1.039)
$NO_2$	1.016	(0.998, 1.033)
$NO_x$	1.012	(0.99, 1.028)
SO <sub>2</sub>	1.009	(0.993, 1.024)
Composite	1.026	(1.004, 1.046)

Summary of covariate and parameter estimates

Summary of covariates

	RR	95% CI
JSA	1.206	(1.195,1.219)
Houseprice	0.945	(0.924,0.969)

Summary of model parameters

	Median		97.5%
$\sigma^2$	0.0349	0.032	0.0379
$\rho$	0.9552	0.9242	0.9766
$\alpha$	0.0349 0.9552 0.8456	0.8179	0.8721

# Conclusions

- New model for air pollution and health that is effective in capturing residual structure in space and time
- The model was applied to a large data set for London, an analysis that is the first of its kind.
- For the London data, tangible improvement in fit and estimation if space-time structure acknowledged
- The new results show that pollution is still a significant factor in respiratory ill health, despite long-term improvements in urban air quality.
- Model is fast to fit.

## Further work

Further development of composite pollution measures.

- Interpretation presently unclear
- Relax assumption that each pollutant contributes equally

Spatio-temporal residual structure not likely to be really smooth.

- Much work in the spatial domain eg. Lu and Carlin (2007); Lee and Mitchell, (2012) but little or none in the space-time setting.
- Approaches involve treating the adjacency structure as not fixed - computationally hard in space, more so in space-time.

## References

Besag, J., J. York, and A. Mollie (1991). Bayesian image restoration with two applications in spatial statistics. *Annals of the Institute of Statistics and Mathematics* 

Leroux, B., X. Lei, and N. Breslow. (1999). Estimation of disease rates in small areas: A new mixed model for spatial dependence, *Chapter Statistical Models in Epidemiology, the Environment and Clinical Trials, Springer-Verlag, New York.* 

Lu, H. and B. Carlin (2005). Bayesian Areal Wombling for Geographical Boundary Analysis. *Geographical Analysis*.

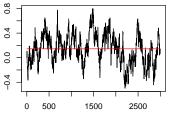
Miller., B. G. (2010). Report on estimation of mortality impacts of particulate air pollution in london. *Institute of Occupational Medicine*.

Lee, D. and R. Mitchell (2012). Boundary detection in disease mapping studies. *Biostatistics*.

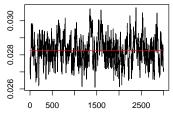
## Thanks for your attention



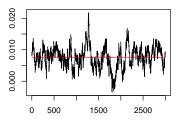
## Trace plots: Fixed effects

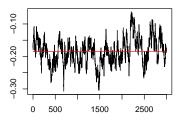






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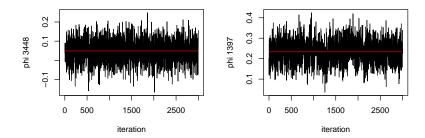


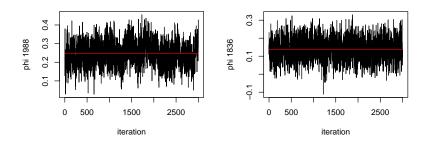


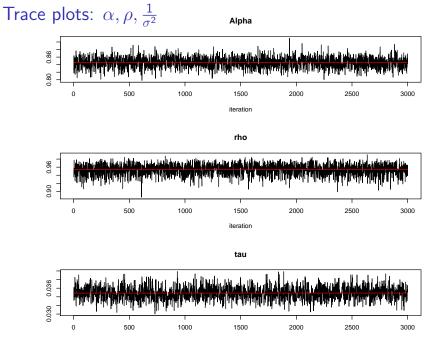


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## Trace plots: $\phi$







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